



# Article Stochastic Modeling of the Al Hoceima (Morocco) Aftershock Sequences of 1994, 2004 and 2016

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Abstract: The three aftershock sequences that occurred in Al Hoceima, Morocco, in May 1994 (Mw 6.0), February 2004 (Mw 6.4) and January 2016 (Mw 6.3) were stochastically modeled to investigate their temporal and energetic behavior. A form of the restricted trigger model known as the restricted epidemic type aftershock sequence (RETAS) was used for the temporal analysis of the selected series. The best-determined fit models for each sequence differ based on the Akaike information criteria. The revealed discrepancies suggest that, although the activated fault systems are close (within 10 to 20 km), their stress regimes change and shift across each series. In addition, a stochastic model was presented to study the strain release following a specific strong earthquake. This model was constructed using a compound Poisson process and depicted the progression of the strain release during the aftershock sequence. The proposed model was then applied to the data. After the RETAS model was used to evaluate the behavior of the aftershock decay rate, the best-fit model was obtained and integrated into the strain-release stochastic analysis. By detecting the potential disparities between the observed data and model, the applied stochastic model of strain release allows for a more comprehensive examination. Furthermore, comparing the observed and expected cumulative energy release numbers revealed some variations at the start of all three sequences. This demonstrates that significant aftershock clusters occur more frequently shortly after the mainshock at the start of the sequence rather than if they are assumed to occur randomly.

Keywords: point process modeling; RETAS model; aftershock energy release; Al Hoceima; Morocco

## 1. Introduction

Seismic events can be classified into three main types based on their distribution over time [1]: (1) mainshock followed by a number of aftershocks decreasing in frequency, (2) slow build-up of seismicity leading to a type (1) sequence and (3) gradual increase and decay of seismicity without a distinct mainshock (seismic swarm), which occurs in areas with complex tectonic structures.

The decrease in aftershock occurrences caused by a strong earthquake can be studied using a wide range of methods, according to [2]. The Omori law model is the most typically adopted model [3], which [4] adapted into the modified Omori formula (MOF) by assuming that the fluctuation of the stress field of the mainshock initiates all the events in the sequence.

The trigger events are conditionally independent and follow a non-stationary stochastic Poisson process. Considering the complex behavior of some earthquake series, particularly in the presence of secondary events, [5] introduced the epidemic type aftershock



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**Copyright:** © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). sequence (ETAS) model, by increasing the capacity to generate secondary events for each event in the sequence. There are several triggering models between these two limit situations: the MOF and ETAS models being two of them [6,7]. The RETAS model [8] was developed by applying the principle of Bath's law [9,10] to the subsequences caused by principal events, such as the mainshock. It is worth mentioning that the magnitude difference between the mainshock and the strongest aftershock is commonly considered to be constant, ranging between 1.2 and 1.4 on average, depending on the criterion [11], although with a lot of variability between individual aftershock sequences [12].

In this study, we focus on three sequences of type (1) designated as the aftershock sequences of Al Hoceima 1994, 2004 and 2016, occurring near the city of Al Hoceima, in Morocco (Figures 1 and 2). The multifractal properties of these sequences have already been investigated in the framework of the spatial modeling of many seismic series in the Ibero–Maghrebian region [13]; however, even more temporal, energy and stress evaluations are required. Therefore, this study aims to examine three aftershock series using stochastic modeling.



**Figure 1.** Geologic sketch map of the Betic-Rif region and location of the Al Hoceima study area. Bk: Bokoya Massif; Ra: Ras Afrou; Rt: Tas Tarf; Tf: Tres Forcas Cape. Internal Zone includes Sebtides and Ghomarides in the Rif and Nevado-Filabrides, Alpujarrides and Malaguides in the Betics.



**Figure 2.** Seismicity recorded by the Spanish IGN included in the 1994, 2004 and 2016 seismic series from magnitude 2.0. Main tectonic features are displayed.

This contribution includes the first section, which describes the regional geological context. In the second section, as suggested by [14], the three aftershock sequences are analyzed, and the stress regime in each series is comprehensively described. The Gutenberg–Richter relationship analysis, performed in a later section, attempts to derive reliable threshold magnitude values and *b*-value estimates for each sequence. A stochastic point process modeling analysis was performed in the previous two sections. As described previously, the aftershock decay rate was comprehensively studied using the RETAS model. The identified best-fit model was then integrated into a stochastic analysis of strain release. A comparison between the real values of the cumulative energy release and the expected modeled values is also examined and addressed.

## 2. Geological Setting Overview

The Rif, along with the Betics, forms the westernmost alpine ranges of the Mediterranean Sea and are linked by the Gibraltar Arc. The Alboran Sea is in the center (Figures 1 and 2). The central-south region of the Alboran Sea and the eastern Rif Cordillera belong to the seismically active area of Al Hoceima. The Internal and External Zones separated by the Flysch Units tectonically constitute the Rif Cordillera. In addition, several late intramontane Neogene–Quaternary sedimentary basins emerged, some of which were linked to the Alboran Sea, forming the largest basin of the orogen [15].

The Internal Zone comprises Sebtide and Ghomaride superposed tectonic complexes formed by Paleozoic, Mesozoic and Cenozoic rocks, which are strongly affected by the Alpine Orogeny and have their equivalents in the Betics, called Alpujarride and Malaguide. Some of these complexes have undergone metamorphism and have been thrust over the Flysch Units and the External Zone in the Rif. The Flysch Units, which are mostly Tertiary sedimentary rocks with some locally ultrabasic rocks, constitute the sedimentary basin that separates the Internal and External Zones and is underlain by oceanic crust. They thrust southward across the External Zone, which is formed by Mesozoic and Cenozoic sedimentary rocks that are mostly unmetamorphosed or, in some cases, have a low degree of metamorphism.

The Alboran Sea is primarily formed by Neogene and Quaternary sediments deposited on a basement that corresponds to the Internal Zone complexes [16]. Furthermore, Neogene to Quaternary volcanic rocks can be found in the central-eastern Alboran Sea and eastern Rif and Betics.

The main alpine deformations in this area occurred throughout the Oligocene and Miocene and continue into the present. The earliest stages of deformation were partially simultaneous with the process of western migration of the Betic and Rif Internal Zones, coinciding with the opening of the Alboran Sea and forming the Gibraltar Arc during a period of severe weakening of the continental crust. The new Alboran marine area corresponds to the western end of the Algero–Provençal Basin, which began to open at the end of the Oligocene [17], forming a new oceanic floor. In the Alboran Sea, the continental crust was markedly weaker and situated on the new oceanic floor on its eastern border.

Subduction processes, combined with NNW–SSE convergence and regional compression of the Iberian and Nubian plates [18–20], resulted in significant deformation in the northern and southern borders of the earliest Alboran Sea. These processes developed the Gibraltar Arc, with the uplift of the Rif and Betic Cordillera, which were radially deformed [21] around the Alboran Sea and undergoing a regional E–W compression.

Later, from the late Miocene, when the opening was nearly at its end ceasing the E–W compression, the general NNW–SSE compression was completely re-established [22]. The region then began to undergo folding (e.g., the Alboran Ridge anticline) and faulting (e.g., the Al Idrisi, Yusuf, Carboneras, Averroes, Jebha and Nekor faults, in addition to other minor faults).

Since the Miocene, important NNE–SSW sinistral strike-slip fault systems crossing the Alboran Sea (Trans Alboran shear zone; [23]) were formed, such as the Carboneras Fault, coming from Almería, Spain and the Nekor and Jebha faults, the last being renowned inland [24]. Later, the Al Idrisi and other conjugated NW–SE faults developed (e.g., the Averroes Fault), as well as some E–W faults and thrusts with a general ENE–SWS strike. The upper Neogene and Quaternary sediments of the central and eastern Alboran Sea are affected by these deformations, whereas the western Alboran Sea undergoes mud volcanic tectonics [25]. This regional geodynamic setting continues into the present.

Almost all the domains mentioned are present in the study area. The Internal Zone forms the inland Bokoya Massif, between Al Hoceima and Melilla to the west, as well as a smaller outcrop, Ras Afrou, on the coast between Al Hoceima and Melilla. The Flysch Units and External Zones, as well as several Neogene–Quaternary basins, appear south of the Internal Zones. Furthermore, Miocene volcanic rocks comprise most of the Raf Tarf and Tres Forcas capes. There are also significant faults (Figures 1 and 2), the most notable of which are located around the Nekor Basin limits, east of Al Hoceima, whose directions range from nearly N–S to NE–SW (e.g., the Trougout Fault, separating Ras Tarf volcanic rocks from Al Hoceima Bay). Some of these faults remain active offshore.

Evidently, all existing seismogenic faults, both onshore and offshore, are unknown, particularly because some are in the early stages of development [26].

#### 3. Aftershock Sequences Description

The Alboran region, specifically the Al Hoceima region, has been the subject of numerous tectonic studies, including those by [27–31]. Owing to its location in the complex border zone between the Eurasian and Nubian plates, near the border between the eastern Rif Cordillera tip and the Alboran Sea, the Al Hoceima region is known to be the most seismically active sector in northern Morocco. Furthermore, because of its strong seismicity, it is one of the most seismically active sites in the western Mediterranean region.

The seismic database of the Spanish Instituto Geográfico Nacional (IGN) was used to assemble the data for this study, with no further processing or parameters from other local or regional agencies. This was performed to keep the database as homogeneous as possible, working in all cases with m<sub>b</sub> magnitude.

The 1994 earthquake sequence (Figure 2) began on 26 May 1994, with a strong earthquake of magnitude Mw 6.0, which struck the coastal region near Al Hoceima. This event had a strong impact on the studied region of Al Hoceima [32]. The maximum felt intensity was VIII–IX (EMS-98), indicating an extended NNE–SSW corridor that accounted for over 80% of the damage reported. The magnitude of the earthquake was revised to Md 5.7 (Moroccan Scientific Institute), and the epicenter was relocated north of Al Hoceima at a focal depth of 13 km. According to [32], the distribution of aftershock epicenters in Figure 2 is largely scattered along a NNE–SSW trending cluster over an almost vertical plane.

Another seismic series struck the region on 24 February 2004, with a damaging mainshock of Mw 6.4 (Figure 2). This event occurred on land and caused severe damage. Ref. [33] estimated the maximum perceptible intensity around XI (EMS-98). In Al Hoceima and the surrounding area, nearly 630 people died, 926 were injured, and nearly 15,000 were left homeless. Ref. [34] relocated the 2004 sequence. According to [35] and other authors, the series epicenter occurred on a NE–SW trending strike-slip fault, while the presence of a NW–SE fault with conjugate NE–SW branches cannot be ruled out [34].

The third examined sequence (Figure 2) is linked to the major event on 25 January 2016 (Mw 6.3), whereas it is possible that the series began on January 21 with an event of Mw 5.1. Following these events, a major earthquake series with decreasing activity occurred in 2016 and 2017 [33,35,36]. The major event, with a maximum intensity of VI–VII, was felt over the Alboran area, particularly in Melilla, Spain, on the northern African coast, where extensive damage was reported, as well as in Al Hoceima [33,37].

The Spanish IGN found two distinct epicenter clusters in 2016, each with distinct tendencies. The first one is aligned NNE–SSW changing to N–S, with dominant strike-slip focal mechanisms, while the second one is to the northeast of the first, with a rounded shape and a dominant reverse focal mechanism solution. A NNE–SSW subvertical fault, roughly parallel to the elongation of the alignment and displaced west of the Al Idrisi Fault, is linked to the main NNE–SSW alignment [26].

Figure 3 shows the number of events per day for the selected series in the 150 days after the mainshock, while Table 1 provides the number of events in each series, the minimum recorded magnitude and other computed parameters, which will be discussed later. The 1994 Al Hoceima sequence had 263 recorded events with a magnitude above 2.0, which occurred until December 1994 (some early events have been included in the series); the 2004 sequence had 969 recorded events with a magnitude above 1.5, which occurred until February 2005; and the 2016 Al Hoceima sequence had 2577 recorded events with a magnitude above 0.8, which occurred until August 2016.



**Figure 3.** Temporal evolution of the studied series (magnitudes above 2.0). Starting of the x-axis does not always correspond to the occurrence of the main event.

**Table 1.** Number of events (n), minimum recorded magnitude ( $M_{min}$ ), computed threshold magnitude ( $m_c$ ) using the Maxc method and a and b-parameters.

Sequence	n	M <sub>min</sub>	m <sub>c</sub>	$a\pm \sigma$	$b\pm \sigma$	Sequence
1994	263	2.0	2.8	$5.05\pm0.02$	$1.01\pm0.07$	1994
2004	969	1.5	3.4	$6.56\pm0.02$	$1.14\pm0.05$	2004
2016	2577	0.8	2.0	$4.86\pm0.03$	$0.82\pm0.02$	2016

For the 1994 and 2004 sequences, Figure 3 reveals a direct decrease in the number of events per day over time, whereas the trend for the 2016 series is more complex, apparently because of the complexity of the rupture(s). There were several different phases in the late aftershock sequence. The recorded occurrences followed a deformation band with two unambiguous alignments with widths of less than 10 to 20 km for the first 30 days [26]. The main alignment, which is moved 5–10 km westward, appears to be spatially associated with the Al Idrisi Fault. A decrease in the seismic activity rate was observed over the next 30 days, resulting in an increase in the width of both alignments, which reached 10–20 km. A clear decrease in activity rate was observed at least 60 days after the mainshock, affecting a wider area than 15–25 km in width as [26] indicated.

# 4. Seismic Series Stress Regime

The IGN database, international agencies and numerous studies aimed at analyzing these seismic series were used to obtain earthquake focal mechanisms for the selected series. Figure 4 shows earthquakes with magnitudes greater than 3.0, from the 1994 Al Hoceima sequence, as well as the focal mechanism of aftershock events computed by [32] (Figure 4A); different solutions for the focal mechanism of the main shock (Figure 4B) from the Global Centroid Moment Tensor (GCMT), International Seismological Centre (ISC); and specific works by [38–41] (for both a pre-event and the main quake) and [29].



**Figure 4.** (**A**) 1994 seismic sequence showing earthquakes with magnitude above 3.0 and computed focal mechanisms for aftershocks [32]. (**B**) Different focal mechanism solutions for the mainquake. GCMT: Global Centroid Moment Tensor; T&al.: [40]; M&R: [38]; ISC: International Seismological Centre; B&B-I and B&B-II: [41], for a pre-event and the main quake, respectively; B&al.: [29]; EA&al.: [32,39].

Figure 5A depicts the distribution of events with magnitudes greater than 3.0, which were included in the 2004 series, as well as the estimated focal mechanisms of the largest events. In addition, Figure 5B shows the earthquakes and computed focal mechanisms of the so-called "eastern cluster" from [34].



**Figure 5.** (**A**) 2004 seismic sequence showing earthquakes with magnitude above 3.0 and computed focal mechanisms for the biggest events. (**B**) Earthquakes with magnitude above 1.0 and focal mechanism solutions for the biggest events (m<sub>b</sub> 1.9–3.0) of the "eastern cluster" studied by [34].

Figure 6 shows events with magnitudes more than 3.0 as well as estimated focal mechanism solutions for the 2016 Al Hoceima sequence.



**Figure 6.** 2016 seismic sequence showing earthquakes with magnitude above 3.0 and computed focal mechanisms.

The stress pattern from the inversion of the available focal mechanism data was used to characterize the three seismic series. It is worth noting that for the inversion process, it was not necessary to select between two available nodal planes. To estimate the different parameters of the reduced stress tensor, we used the improved right dihedron method [42] combined with the iterative rotational dihedron method [43]. Our aim was to calculate the four parameters of the reduced stress tensor,  $\sigma_1$ ,  $\sigma_2$ ,  $\sigma_3$  and the stress ratio,  $R = (\sigma_2 - \sigma_3)/(\sigma_1 - \sigma_3)$ .

According to [19], this method allows for estimating previous parameters and the extraction of filtered focal mechanism data by deleting nodal planes that are incompatible with the average stress regime. The compatible focal mechanisms and the calculated stress tensor produced at this point were then employed as the direct starting point for the rotational optimization technique.

The iterative grid-search rotational optimization process is based on a controlled grid search of the stress tensor using the Win-TensorTM code to reduce the so-called misfit (F5) [44]. According to [45], the nodal plane best explained by the stress tensor was chosen as the actual fault plane from the two planes of the focal mechanism. Consequently, the final inversion examines the focal planes that a uniform stress field best fits [46]. After the ultimate optimization, the omitted focal planes must be reconsidered without modifying the stress tensor. If this is the case, the data are re-entered into the database, the stress tensor is re-optimized, and the software runs another check for the rejected data.

This method was used to analyze the focal mechanism data for each seismic sequence. Table 2 and Figure 7 show the results of the stress inversion. Several authors (e.g., [19,44,47–49]) have used a similar approach to investigate the stress regime in other regions.

Sequence	$\sigma_1$	$\sigma_2$	$\sigma_3$	R	F5
1994 main event	327° N/13°	$226^{\circ} \text{ N}/40^{\circ}$	$072^{\circ} \text{ N}/47^{\circ}$	0.19	7.1
1994 aftershocks	$132^{\circ} \text{ N}/44^{\circ}$	349° N/39°	$242^{\circ} \text{ N}/20^{\circ}$	0.38	3.5
2004 all, without 'eastern' cluster	$330^{\circ} \text{ N}/25^{\circ}$	$149^{\circ} \text{ N}/65^{\circ}$	$239^{\circ} \text{ N}/00^{\circ}$	0.59	2.6
2004 'eastern' cluster	$020^{\circ} \text{ N}/79^{\circ}$	$141^{\circ} \text{ N}/06^{\circ}$	$232^{\circ} \text{ N}/09^{\circ}$	0.82	5.6
2016 main cluster	332° N/06°	224° N/72°	$064^{\circ} \text{ N}/17^{\circ}$	0.74	1.6
2016 secondary cluster	330° N/23°	$062^{\circ} \text{ N}/04^{\circ}$	$161^{\circ} \mathrm{N}/67^{\circ}$	0.49	3.3

Table 2. Stress regime for the cases considered in Figure 7.

The variability in the stresses is highlighted in Figure 7 and Table 2. The different solutions of the mainshock, when considered together, and those determined for the aftershocks were not very distinct in the 1994 seismic series. A prolate stress ellipsoid with a near NW–SE horizontal (compressive stress regime) and roughly similar  $\sigma_2$  and  $\sigma_3$  values appeared in the mainshock. The slope then turns southeast, highlighting the NE–SW expansion with the horizontal  $\sigma_3$  being more noticeable (extensive stress regime).

The seismic series from 2004 appears to be more uniform, with oblate stress ellipsoids and a noticeable NE–SW extension trend. While the main cluster shows NW–SE subhorizontal  $\sigma_1$  and NE–SW sub-horizontal  $\sigma_3$  values, clearly indicating a strike-slip stress regime, in the "eastern cluster" [34], a cluster with a few low-energy earthquakes and  $\sigma_1$ and  $\sigma_2$  values becoming closer, and a NE–SW horizontal extension is dominant (extensional stress regime).

In the 2016 sequence, the computed stress regime for the main cluster agrees with that obtained in the 2004 main cluster (strike-slip stress regime); however, there is a noticeable increase in the axial ratio, where  $\sigma_1$  and  $\sigma_2$  magnitudes are closer, highlighting the well-defined NE–SW sub horizontal extensive stresses. In contrast, the secondary cluster is dominated by NW–SE compressive  $\sigma_1$  stresses and a subvertical  $\sigma_3$ , suggesting thrusting (compressional stress regime). Furthermore, the secondary cluster stresses are like those obtained for the mainshock of 1994, which had a primary NNW–SSE  $\sigma_1$  odd axis.

The maximum compressive horizontal stress (SHmax) in all solutions was between 142° N and 153° N, which is consistent with previous results for regional stresses (e.g., [18–20,50]). However, the extension directions are compatible with the Internal Zone movement of the Betic–Rif toward the southwest [26,51–53].







**Figure 7.** Stress regime computed from focal mechanisms. (**A**) For the 1994 mainshock, (**B**) for the 1994 aftershock sequence, (**C**) for the 2004 aftershock sequence, (**D**) for the 'eastern cluster' considered by [34], (**E**) for the 2016 'main cluster' and (**F**) for the 2016 'secondary cluster'.

# 5. Magnitude–Frequency Relationships

The Gutenberg–Richter recurrence relationship [54,55] is a frequently used approach for quantifying seismic activity in each region and has been shown to apply aftershock frequency–magnitude data. The equation is a reasonable approximation of the frequency– magnitude statistics that describe the correlation between earthquake occurrence frequency and magnitude

$$\log_{10} N \ (\ge m) = a - bm \qquad m \ge m_c \tag{1}$$

where  $N (\geq m)$  is the number of events with magnitudes greater than or equal to m. For the estimation of both the a and *b*-values, it is widely recommended to use a complete dataset for all ranges of magnitude.

The threshold magnitude parameter  $m_c$ , is typically determined using one of two methods: a network-based [56,57] or catalog-based approach [56,57]. The first group uses the day-to-night ratio to calculate the earthquake frequency, if noise reduces the detection threshold at night [58,59]. The second set of approaches assumes that earthquake production is self-similar, allowing us to use a power law or the Gutenberg–Richter relationship to construct earthquake frequency–magnitude distributions. The most applied methods in this group reviewed by [60,61] are the maximum curvature (Maxc) method [62], entire magnitude range (EMR) method [60,63], median-based analysis of the segment slope (Mbase) [64], determination of *b*-value instability [65] and goodness-of-fit test (Gft) by [62], which was later modified by [66].

The threshold magnitude for each of the three sequences under consideration was thoroughly estimated in the current study. Although both the maximum curvature (Maxc), Gft, Gft at levels of 5% and 10% (Gft5% and Gft10%), and modified goodness of fit (mGft) methods were first investigated, the maximum curvature method produced better results, i.e., a better fit (Table 1). Threshold magnitude values of 2.8, 3.4 and 2.0 were obtained for the 1994, 2004 and 2016 aftershock sequences, respectively. The threshold value fluctuated over time, which was also studied. This temporal fluctuation is depicted in Figure 8 and was computed using a sliding-window method.

Prior to each new computation, a window of 20 events was shifted by five events [67]. The middle time of the associated window is supplied to each new threshold magnitude value. The window length was chosen as the balance between the need for temporal resolution and smoothness. Multiple tests were conducted first, changing the number of occurrences per window and the shift; however, neither aspect had a significant influence on the definition. The maximum likelihood approach was used and is considered one of the most reliable approaches for computing the *b*-value among the available methods. The estimator by [68], given below, was used to calculate the *b*-value:

$$b = \frac{\log_{10}e}{\langle m \rangle - m_c} \tag{2}$$

 $\langle m \rangle$  being the average value of the magnitude. An estimate of the standard deviation  $\sigma_b$  is obtained using the [69] relation, given as follows:

$$\sigma_b = 2.30 \ b^2 \sqrt{\frac{\sum_{i=1}^N (m_i - m)^2}{N(N - 1)}} \tag{3}$$

According to [69,70], even when the *b*-value varies in time and/or space, this relationship provides a reliable approximation of *b*-uncertainty. When utilizing rounded magnitudes, the estimator by [68] is inaccurate but agrees with the modified distribution of [71]. Then, the improved estimator by [72], given below, is used to determine the maximum likelihood *b*-value:

$$b = \frac{\log_{10}e}{\langle m \rangle - \left(m_c - \frac{\Delta m}{2}\right)} \tag{4}$$

The estimated *b*-value appraisal for the 1994, 2004 and 2016 aftershock sequences, using the maximum likelihood estimator by [72], with a bin width of 0.1 units, are  $1.01 \pm 0.07$ ,  $1.14 \pm 0.05$  and  $0.82 \pm 0.02$ , respectively (Table 1). Figure 9 shows the power law distributions for the three sequences as fitted by a straight line. The greatest magnitudes, corresponding to the mainshock, are not explained by the Gutenberg–Richter relationship and are thus regarded as outliers in all series. It is worth mentioning that the estimated



maximum likelihood parameters are closely linked to and impacted by the previously determined threshold magnitude.

Figure 8. Temporal evolution of the threshold magnitude for the three sequences.



**Figure 9.** The cumulative number of earthquakes vs. magnitude (black points and linear fits) and the non-cumulative number of earthquakes (open points) for the three sequences. The threshold magnitude is highlighted.

#### 6. Temporal Stochastic Modeling for the Al Hoceima Sequences

This section focuses on the point process modeling of the studied aftershock sequences. Many approaches for modeling this gradual decrease in aftershock frequency have been presented in the literature. The most commonly used model is the Omori law [3], which was improved by [4] into the modified Omori formula (MOF). According to [4], the decay rate of the aftershock per unit time is given as follows:

$$n(t) = \frac{K}{\left(t+c\right)^p} \tag{5}$$

where *t* is the time after the occurrence of the mainshock, and the parameter *K* is related to the mainshock event and threshold magnitude. The *c* parameter is a debatable number [73,74], with the early stages of imprecise detection of low events in the sequence having a significant effect [75]. Finally, the *p* parameter is a decay constant, and it is quite likely the most important parameter for understanding the behavior of the sequence. The *p*-value varies from sequence to sequence and typically ranges from 0.5 to 1.8 [73]. This variation may be related to tectonic activity in the area. However, the elements that influence the *p*-value remain unclear [76,77].

The quantity n(t) allows us to connect with the point process model by considering it a conditional intensity.

K

$$u(t) \approx \check{}(t), \tag{6}$$

bearing in mind that [5]:

$$P \{an \text{ event occurs in } (t, t+dt) | \mathfrak{T}_t\} = \check{}(t|\mathfrak{T}_t) dt + o(dt), \tag{7}$$

where  $\Im_t$  denotes the internal history of the occurrence process at time t,  $\check{}(t|\Im_t)$  is the conditional intensity function [5,78], and o(dt), in the Landau notation, is a function of a lower order than the function  $\lambda$ , i.e., o(dt) being negligible. The MOF model includes only the mainshock occurrence time because it is based on the concept that the entire relaxation process is driven by the stress changes induced by the mainshock alone. The aftershocks are conditionally independent and follow a non-stationary Poisson process.

The MOF model fits data for simple aftershock sequences well; however, secondary clustering is common when there are secondary aftershocks triggered by strong earthquakes in a sequence. The authors in [5] argued that aftershock clustering is a self-similar process in which all aftershocks might induce other aftershocks, with a triggering capacity dependent on their magnitudes, because of these complex situations involving one or more secondary events.

The model was named ETAS, and its conditional intensity function is given as follows [79]:

$$\tilde{}(t|\mathfrak{T}_{t}) = \mu + \sum_{i; \ t_{i} < t} \frac{k_{0} e^{\alpha(m_{i} - m_{c})}}{(t - t_{i} + c)^{p}}$$
(8)

where  $\mu$  is the background seismicity rate. The internal history  $\Im_t$  includes the time occurrence  $t_i$  (in days after the mainshock) and magnitude events  $m_i$  of all the events occurring before time t. The summation includes all events with occurrence times  $t_i$  and magnitudes equal to or greater than the lower cut-off  $m_c$ . The c and p model parameters were the same as those used in the MOF model. Furthermore,  $k_0$  is a parameter that affects total aftershock productivity and is common to all aftershocks.

According to [80], every term of the summation in Equation (8) indicates the contribution of a prior event to the occurrence probability of subsequent events at time t. The exponential term is controlled by two factors: (a) the temporal rate decrease, as presented by the MOF model and (b) the exponential term chosen because the logarithm of the aftershock area and the magnitude of the mainshock present a linear relationship [81]. Parameter  $\alpha$  measures the effect of the magnitude of the production of 'children' events, also called "descendants".

The MOF and ETAS models, as expressed by Equations (5) and (8), respectively, present two limited cases for modeling the temporal distribution of an aftershock sequence. Ref. [2] proposed a similar model, the RETAS model, based on the assumption that not all events in a series; however, only aftershocks with magnitudes greater than or equal to a threshold value, can directly cause the aftershocks of "descendants". This model allows for the inclusion of all potential models between the two limit cases of the MOF and ETAS models, resulting in the conditional intensity expressed as follows:

$$\tilde{(t|\mathfrak{F}_{t})} = \mu + \sum_{\substack{i; \ t_{i} < t \\ m_{i} \ge M_{tr}}} \frac{k_{0} e^{\alpha(m_{i} - m_{c})}}{(t - t_{i} + c)^{p}}$$
(9)

It is worth noting that the RETAS model developed by [8] is based on Bath's law, which states that the difference between mainshock magnitude and the strongest aftershock magnitude, ranges between 1.2 to 1.4 units. According to this relationship, ref. [8] argued that the difference between the weaker primary event and the weakest event in the after-shock sequence must be at least 1.2 units, by applying this principle to the subsequences generated by the primary trigger model.

Furthermore, the RETAS model has the advantage of examining all potential models between the MOF and ETAS models because the triggering magnitude ranges from the threshold magnitude to the mainshock magnitude. The Akaike information criterion [82], abbreviated as AIC, was used to choose the best-fit model in our case, with the lower AIC value. This is expressed as follows:

$$AIC = -2 \max \log L(\theta; 0, T) + 2k^{"}$$
(10)

where  $k^{"}$  represents the number of parameters of the model, and *log L* is the logarithm of the likelihood function, given as follows:

$$\log L(\theta; 0, T) = \sum_{i=1}^{N} \log_{10} \tilde{\theta}(t_i | \mathfrak{T}_{t_i}) - \int_{0}^{T} \tilde{\theta}(s | \mathfrak{T}_s) ds$$
(11)

In the previous Equation (12), N is the number of earthquakes with magnitudes greater than or equal to  $m_c$ , occurring at times  $t_j$  (j = 1, 2, ..., N), during [0, T]. Due to the features of the described model, it was used in this study to investigate three aftershock sequences. Table 1 lists the magnitude thresholds used in this study.

As a result of the RETAS model analysis, Figure 10 displays the AIC value versus the triggering magnitude value. Analysis of this curve reveals the aftershock clustering type that is most common in the sequence. The AIC parameter was calculated for all the potential models by varying the triggering magnitude from the threshold magnitude to the magnitude of the mainshock. The best-fit model shows the AIC minimum values in Table 3 as well as the triggering magnitude and model parameters.



Figure 10. The AIC values of the RETAS model vs. the triggering magnitude for the three sequences.

Table 3. The computed pa	arameters for the RETAS model.
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Sequence	Mtr	Best Model	AIC	K <sub>0</sub>	α	с	р
1994	3.9	RETAS	-12.4	1.727	0.0630	0.024	0.909
2004	3.4	ETAS	-1222.5	3.426	0.0035	0.082	1.070
2016	2.0	ETAS	-7209.0	1.477	0.0380	0.039	1.183

When examining the AIC values vs.  $M_{tr}$  curves (Figure 10), we can see that the best-fit model for the minimal AIC value was found to be for a  $M_{tr}$  value of 3.9 for the 1994 sequence. For the 2004 sequence, an  $M_{tr}$  value of 3.4 is the best-fit model obtained and, for the 2016 sequence, a magnitude of 2.0. The best-fit model became an ETAS model when the triggering magnitude  $M_{tr}$  coincided with the completeness threshold magnitude  $m_c$  in the last two  $M_{tr}$ . Analyzing the AIC vs.  $M_{tr}$  trigger magnitude curve for the 2016 sequence,

a monotonic trend exists, resulting in a continuous increase in AIC values, the lowest of which is for  $M_{tr} = m_c$ .

According to the model presented in Figure 11, the estimated best model parameters shown in Table 3 were used to compute the expected cumulative number of earthquakes and error bounds. The computed values were compared with the observed cumulative number of earthquakes. Figure 11 shows that the estimated model matches the observed data well for all three sequences, with the observed values remaining inside the error boundaries throughout the whole sequence.



**Figure 11.** The cumulative number of events above magnitude  $m_c$  for the three sequences. Circles: observed values. Lines: mean  $\pm \sigma$  fitted model. The magnitudes of the events included in the sequences are also depicted. Inset stand out the first ten days of the series.

#### 7. Stochastic Model for the Energy Release

In the literature, several attempts have been made to improve earthquake time models by including data on other characteristics associated with event occurrences, such as space-time models or those that relate event occurrence time to a size value of the event (intensity, magnitude or energy). In this context, this section aims to investigate the relationship between the occurrence times  $\{T_k, k = 1, ..., n\}$  and released energy  $\{E_k, k = 1, ..., n\}$ , providing information on the size of the events in the studied sequences. It is typically considered that the time of occurrence and scale of an event are unrelated.

As a result, our aim was to create an energetic stochastic model that incorporates these considerations and examines how well it represents the observed data. According to [83], marked point processes have similar counting processes [84,85]. The following is a description of this model. Given a Poisson process {N(t);  $t \ge 0$ } with a rate  $\lambda > 0$ , it is assumed that the time  $T_k$  of each event is linked to a realization, which is a family { $Y_k$ ; k > 0} of independent and identically distributed random variables called marks, with a probability distribution function equal to

$$G(y) = P\{Y_k \le y\} \tag{12}$$

The second requirement is that the random variables are at the same time independent from  $\{N(t); t \ge 0\}$ . Then, according to [84,85], the stochastic model, typically called the compound Poisson process, is defined as follows:

$$Z(t) = \sum_{k=1}^{N(t)} Y_k; \quad t \ge 0$$
(13)

Denoted by  $\mu$  and  $\gamma^2$  are the mean and variance of the marks  $Y_k$ ; the moments of Z(t) are then given as follows:

$$E[Z(t)] = \check{\mu}t \tag{14}$$

$$var[Z(t)] = \check{}\left(\gamma^2 + \mu^2\right)t \tag{15}$$

Consequently, in the case of a series of occurrence times and event sizes, the compound Poisson process can be used as a model process for random behavior. In the most general treatment of a compound point process, {N(t);  $t \ge 0$ } is an inhomogeneous Poisson process with an intensity function { $\lambda(t)$ ;  $t \ge 0$ }, and marks { $Y_k$ ; k > 0} do not have to form an independent series of random variables. Equations (14) and (15) must be rewritten in the context of an inhomogeneous Poisson process with a rate that varies with time  $\lambda = \lambda(t)$ 

$$E[Z(t)] = \mu \int_{t_0}^t (s) ds$$
 (16)

$$var[Z(t)] = \left(\gamma^2 + \mu^2\right) \int_{t_0}^t \check{}(s)ds$$
(17)

The following relationship (18) between the log of the released energy and the magnitude is used [55,86] because the most common way of describing the size of an earthquake is by its magnitude.

$$\log_{10}\sqrt{E} = 2.4 + 0.75 \,M \tag{18}$$

The relationship described by this Equation (18), according to [86,87], is consistent with what is expected theoretically for classical crack models with a constant stress drop. This generalized energy–magnitude scaling equation works for various magnitude ranges.

As the cumulative released energy is a physical quantity, it is helpful to consider the series of released energies in the compound Poisson process scheme. It is worth noting that this method has several limitations. First, the energy is determined by the wide dynamic

range of the released energy, as well as an instrumentation earthquake record. Thus, to reduce this variability, Equation (18) is transformed as follows:

$$E_k^{tr} = \frac{E_k}{E_0} = \frac{10^{4.8+1.5} M_k}{10^{4.8+1.5} M_0} = 10^{1.5 (M_k - M_0)}$$
(19)

where  $E_k$  and  $M_k$  correspond to the energy and magnitude of the *k*-th event, respectively, and  $E_0$  and  $M_0$  are the energy and magnitude of the first event considered, respectively.  $E_k^{tr}$  is the Benioff's energy.

The approach described above was used to analyze the considered aftershock sequences. Considering the compound Poisson process, given as follows:

$$Z(t) = \sum_{k=1}^{N(t)} E_k^{tr}$$
(20)

Equations (16) and (17) were used to derive the estimation of the mean E[Z(t)] and the variance var[Z(t)] for each series, considering the best-fit model derived and analyzed in the previous section. Figure 12 displays the expected cumulative energy and confidence bounds, mean plus/minus the standard deviation, according to the appropriate best-fit model for each aftershock sequence, and this is compared to the computed cumulative energy released in the sequences.



Figure 12. Cont.



**Figure 12.** Cumulative released energy for the three sequences. Circles: observed values. Lines: mean  $\pm \sigma$  fitted model. Magnitudes of the events included in the sequences are also depicted. The inset shows the first ten days of the series.

It should be noted that the stochastic model used follows the two previously quoted assumptions: both the independence of the magnitudes and the independence of the magnitudes with the occurrence time. Figure 12 shows that, for the 1994 sequence, no significant deviation of real data values from the model was observed, which supports the above assumptions. However, deviations of the real data values with the model were observed for the two other aftershock sequences.

For the 2004 sequence, the main deviation was observed at the beginning of the sequence, up to the 30th day, where the clustering of stronger aftershocks was recognized. This reveals that the assumption of the independence of magnitudes with occurrence times is not valid in this case. In addition, another deviation of the observed data out of the error bounds can be observed in the 2016 sequence, where an additional concentration of stronger aftershocks is recognized.

### 8. Discussion

In the present study, the stress pattern analysis revealed that all solutions had a maximum compressive horizontal stress (SHmax) between 142° N and 153° N, which is compatible with previous regional stress data as determined by previous studies. Furthermore, according to these results, the extension directions are clearly compatible with the Betic–Internal Rif's zone movement southwestward.

The results for all the seismic sequences share horizontal stresses, showing a NW–SE compression and a NE–SW extension as a result of the regional setting. Nonetheless, some significant differences can be observed in the axial ratios and local stress regimes because of the series location in the main shear zone crossing the Alboran Sea and the activated structures in the main and secondary clusters. The three series induced a different fault system, also hosting seismicity, from which the sequence began.

Concerning observed differences in the tectonic characteristics of the series, the 1994 sequence is mainly scattered along a NNE-SSW trending cluster over an almost vertical plane [32], the 2004 sequence is mainly scattered along a NE-SW strike slip fault, and the 2016 sequence is initially scattered along a NNE-SSW trending cluster over a subvertical strike-slip fault, then changing to N-S and finally distributed on a second rounded cluster with a dominant reverse focal mechanism solution [26]. The magnitude of completeness values considered in the current study were compared with those derived using Gft and mGft methods.

Figure 13 shows plots of residuals vs. the minimum magnitudes for the three series. The results obtained matched the estimated values derived using the maximum curvature (Maxc) method. All approaches yield a result of 2.8 for the 1994 series. For the 2004 series,

the difference was of the order of 0.1 using mGft, Gft5% and Maxc and 0.2 between Gft10% and Maxc. However, in the case of the 2016 series, the difference varies between 0.2 and 0.3. The results are 1.9 using Gft10%, 2.2 for Gft5%, 2.3 for mGft and 2.0 for Maxc.



**Figure 13.** The obtained residuals vs. magnitude when fitting the observed data to the mGft power law (solid line) and theoretical distribution power law (dashed line). The 5% and 10% residual levels are shown as reference.

As discussed in this study, the RETAS model was used to study the gradual decay of the aftershock frequency based on the triggering magnitude  $M_{tr}$  assumption. The minimum value of the AIC parameter was used as a criterion for selecting the best-fit model. All possible models were estimated by varying the triggering magnitude  $M_{tr}$  from the threshold magnitude  $m_c$  (ETAS model) or  $m_{ms}$  (MOL model) to the mainshock magnitude. The results derived for the 1994 and 2004 series highlight and improve on the previous ones derived by [88]. For instance, for the 2004 aftershock sequence, as in [88], the minimum AIC value shows results for the best ETAS model. Considering that the 2016 sequence began on January 21 with an event of Mw 5.1, as suggested by some authors [26], the ETAS model was obtained as the best-fit model.

Stochastic modeling of the energy revealed that the energy released over the series was outlined well by the proposed model in the 1994 and 2004 series. However, there were minor discrepancies in the 2016 sequence and the computed cumulative released energy did not match the model well. We deduced from the analysis that large aftershock clusters occurred more frequently and quickly after the main shock than they would if they occurred at random. This observation holds true for the Zemmouri, Algeria aftershock series of 2003 [88].

### 9. Conclusions

Stochastic modeling was used to analyze three sequences of events that occurred in Al Hoceima, Morocco, May 1994, February 2004 and January 2016. The analysis of the behavior of the decay rate of the three series, together with the composite stress pattern and the obtained *b*-value, led us to conclude that the 2016 sequence is the most complex of the three series, likely because of two different fault systems being activated—the main one corresponding to a strike-slip stress and the second one to reverse faults.

The released energy analysis allowed us to characterize the occurrence of large aftershocks shortly after the mainshock better than it would if they occurred at random. However, further research is needed to estimate the recurrence period of such large occurrences, as well as the probability of exceeding a specific magnitude shortly after the mainshocks. This type of research could be conducted in other seismically active regions to investigate the behavior of the seismic series occurring in these areas. This could help to understand the characteristics of the earthquake generation process aimed at seismic forecasting studies.

It should be noted that earthquake forecasting is the ultimate challenge for seismologists because it accumulates scientific knowledge about the earthquake occurrence process and is an essential component of any efficient risk-mitigation strategy [89,90]. As stated by different authors (e.g., [80,91]), RETAS models can be applied to forecast the occurrence probability evolution of a certain sequence, providing the possibility to identify the type of clustering in future seismic series. This is an issue also related to, for instance, the fault distribution and possible fault interactions in the studied area.

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